

# Angewandte Statistik UE

V) 1a)  $X_1, \dots, X_n$  Stichprobe;  $X \sim A_\theta$

ges.: (i) FRC-Schranke, (ii) FRC-effizienten Schätzer für  $\theta$ .

(i) Sei  $T$  eine unverzerrte Schätzfkt. für  $\theta$ , d.g.

$$\text{Var}_\theta T \geq \frac{\left[ \frac{\partial}{\partial \theta} \int \theta \right]^2}{n \cdot \mathbb{E}_\theta \left[ \frac{\partial}{\partial \theta} \ln P(X|\theta) \right]^2} =: \text{FRC}$$

$$X \sim A_\theta \Rightarrow P([X=x|\theta]) = \theta^x (1-\theta)^{1-x} \quad (x=0,1) \quad \theta \in (0,1).$$

$$\mathbb{E}_\theta X = \theta, \quad \text{Var}_\theta X = (1-\theta)\theta. \Rightarrow \mathbb{E}_\theta X^2 = \theta.$$

$$\mathbb{E}_\theta \left[ \frac{\partial}{\partial \theta} \ln P(X|\theta) \right]^2 = \sum_{x=0}^1 \left[ \frac{\partial}{\partial \theta} \left( \theta^x (1-\theta)^{1-x} \right) \right]^2 \theta^x (1-\theta)^{1-x}$$

$$= \sum_{x=0}^1 \left[ x \ln \theta + (1-x) \ln(1-\theta) \right]^2 \theta^x (1-\theta)^{1-x}$$

$$= \left( -\frac{1}{1-\theta} \right)^2 (1-\theta) + \frac{1}{\theta^2} \theta = \frac{1}{1-\theta} + \frac{1}{\theta} = \frac{1}{\theta(1-\theta)}$$

$$\Rightarrow \text{FRC} = \frac{\theta(1-\theta)}{n}$$

(ii) Sei  $T = t(X_1, \dots, X_n) = \bar{X}_n$ . ~~NEIN~~

$$\mathbb{E}_\theta(T) = \mathbb{E}_\theta \left( \frac{1}{n} \sum_{i=1}^n X_i \right) = \frac{1}{n} n \mathbb{E}_\theta X = \theta$$

$$\mathbb{E}_\theta T^2 = \mathbb{E}_\theta \left( \frac{1}{n^2} \left( \sum_{i=1}^n X_i \right)^2 \right) = \frac{1}{n^2} \left( \sum_{i=1}^n \mathbb{E} X_i^2 + \sum_{\substack{i,j=1 \\ i \neq j}}^n \underbrace{\mathbb{E} X_i \mathbb{E} X_j}_{(\mathbb{E} X)^2} \right)$$

$$= \frac{1}{n^2} \left( \sum_{i=1}^n X_i^2 + \sum_{\substack{i,j=1 \\ i \neq j}}^n X_i X_j \right)$$

$$= \frac{1}{n^2} \left( n \underbrace{\mathbb{E} X^2}_\theta + n(n-1) \underbrace{(\mathbb{E} X)^2}_{\theta^2} \right) = \frac{1}{n} (\theta + (n-1)\theta^2)$$

$$\Rightarrow \text{Var}_\theta T = \frac{\theta}{n} + \frac{n-1}{n} \theta^2 - \theta^2 = \frac{\theta - \theta^2}{n} = \frac{\theta(1-\theta)}{n} \Rightarrow \text{FRC-effizient}$$

$$\frac{1}{n} \text{Var}_\theta X$$

$$\text{Var}_\theta T = \text{Var}_\theta \bar{X}_n = \text{Var}_\theta \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n^2} \text{Var}_\theta \sum_{i=1}^n X_i = \frac{1}{n^2} n \sum \text{Var}_\theta X_i = \frac{1}{n} \text{Var}_\theta X.$$

b) selbe Diskussion für  $X \sim P_\mu$ .

$$X \sim P_\mu \Rightarrow P[X=x|\mu] = \frac{\mu^x}{x!} e^{-\mu} \quad x \in \mathbb{N}_0, \mu \in (0, \infty)$$

$$\Rightarrow EX = \text{Var } X = \mu \Rightarrow EX^2 = \mu + \mu^2 = \mu(1+\mu)$$

$$(i) E_\mu \left[ \frac{\partial}{\partial \mu} \ln P(X|\mu) \right]^2 = \sum_{n=0}^{\infty} \left[ \frac{\partial}{\partial \mu} \ln \left( \frac{\mu^n}{n!} e^{-\mu} \right) \right]^2 \frac{\mu^n}{n!} e^{-\mu}$$

$$= \sum_{n=0}^{\infty} \left( \frac{n}{\mu} - 1 \right)^2 \frac{\mu^n}{n!} e^{-\mu} = E \left( \frac{X}{\mu} - 1 \right)^2 = \frac{EX^2}{\mu^2} - 2 \frac{EX}{\mu} + 1 = \frac{1+\mu}{\mu} - 1 = \frac{1}{\mu}$$

$$\frac{X^2}{\mu^2} - 2 \frac{X}{\mu} + 1$$

$$\Rightarrow \text{FRC} = \frac{\mu}{n}$$

(ii) Sei  $T := \bar{X}_n$ , d.g.

$$E_\theta T = EX = \mu$$

$$E_\theta T^2 = \frac{1}{n} \left( \underbrace{EX^2}_{\mu + \mu^2} + (n-1) \underbrace{(EX)^2}_{\mu^2} \right) = \frac{\mu}{n} + \mu^2$$

$$\Rightarrow \text{Var } T = \frac{\mu}{n}$$

" "  
" "  
 $\frac{1}{n} \text{Var } X$

2) ZZ:  $E_\theta \left[ \frac{\partial \ln f(x|\theta)}{\partial \theta} \right]^2 = -E_\theta \left[ \frac{\partial^2 \ln f(x|\theta)}{\partial \theta^2} \right]$

$$\int \left[ \frac{\partial \ln f(x|\theta)}{\partial \theta} \right]^2 f(x|\theta) d\lambda(x) + \int \left[ \frac{\partial^2 \ln f(x|\theta)}{\partial \theta^2} \right] f(x|\theta) d\lambda(x) =$$

$$= \int \left( \frac{1}{f(x|\theta)} \frac{\partial f(x|\theta)}{\partial \theta} \right)^2 f(x|\theta) d\lambda(x) + \int \frac{\partial}{\partial \theta} \left[ \frac{1}{f(x|\theta)} \frac{\partial f(x|\theta)}{\partial \theta} \right] f(x|\theta) d\lambda(x)$$

$$= \underbrace{\int \frac{1}{f(x|\theta)} \left( \frac{\partial f(x|\theta)}{\partial \theta} \right)^2 d\lambda(x)}_{=0} + \int \left( -\frac{1}{f(x|\theta)^2} \left( \frac{\partial f(x|\theta)}{\partial \theta} \right)^2 + \frac{1}{f(x|\theta)} \frac{\partial^2 f(x|\theta)}{\partial \theta^2} \right) f(x|\theta) d\lambda(x)$$

$$= \int \frac{\partial^2 f(x|\theta)}{\partial \theta^2} = \frac{\partial^2}{\partial \theta^2} \underbrace{\int f(x|\theta)}_1 = 0.$$

$$10) E_\theta \left[ \frac{\partial \ln P(X|\theta)}{\partial \theta} \right]^2 = \frac{1}{\theta(1-\theta)} \quad E_\theta \left[ \frac{\partial^2 \ln P(X|\theta)}{\partial \theta^2} \right] = -\frac{\theta}{\theta^2} - \frac{1-\theta}{(1-\theta)^2} = -\left( \frac{1}{\theta} + \frac{1}{1-\theta} \right)$$

$$E_\mu \left[ \frac{\partial \ln P(X|\mu)}{\partial \mu} \right]^2 = \frac{1}{\mu} \quad E_\mu \left[ \frac{\partial^2 \ln P(X|\mu)}{\partial \mu^2} \right] = E_\mu \left( -\frac{n}{\mu^2} \right) = -\frac{\mu}{\mu^2} = -\frac{1}{\mu}$$

$\frac{\partial}{\partial \mu} (X-1)$

4)  $X_1, \dots, X_n$  Stichprobe;  $X \sim N(\mu, \sigma^2)$  [ $\mu$  bekannt]

Schätzer für  $\sigma^2$ :

$$S_1 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2; \quad S_2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

a) Schätzer unverzerrt?

1)  $\mathbb{E} S_2 = \sigma^2$  (siehe VO)

2) ~~Wissen:  $\frac{n-1}{n} \frac{S_2}{\sigma^2} = \frac{S_1}{\sigma^2} \sim \chi_n^2$~~

Satz 1.3.3:  $\underbrace{\frac{n}{\sigma^2} S_1}_{=: Y_n} \sim \chi_n^2$ .

$$\Rightarrow \mathbb{E} S_1 = \frac{\sigma^2}{n} \mathbb{E} Y_n = \frac{\sigma^2}{n} n = \sigma^2$$

b) Welcher Schätzer ist effizienter?

1)  $\text{Var } S_2 \stackrel{\text{Ü 3.1, Prop 3}}{=} \frac{2\sigma^4}{n-1}$

2)  $\text{Var } S_1 = \frac{\sigma^4}{n^2} \text{Var } Y_n = \frac{\sigma^4}{n^2} 2n = 2 \frac{\sigma^4}{n}$

$$\Rightarrow \text{Var } S_1 < \text{Var } S_2$$

$\Rightarrow S_1$  effizienter

c) Schätzer konsistent?

erwartungstreu, asymptotisch effizient  $\Rightarrow$  konsistent

d) ZZ:  $S_1$  ist CAN-Schätzfolge

1)  $S_1$  konsistent

2) ZZ:  $\sqrt{n} (S_1 - \sigma^2) \xrightarrow{vt} N(0, \psi(\sigma^2))$

$$K_i := \left( \frac{X_i - \mu}{\sigma} \right)^2 \sim \chi_1^2 \Rightarrow \mathbb{E} K_i = 1, \text{Var } K_i = 2$$

$$S_1 = \sigma^2 \cdot \bar{K}_n = \sigma^2 \cdot \frac{1}{n} \sum_{i=1}^n K_i$$

$$\text{ZGV} \Rightarrow \sqrt{n} \frac{\bar{K}_n - 1}{\sqrt{2}} = \sqrt{n} \frac{\frac{S_1}{\sigma^2} - 1}{\sqrt{2}} = \sqrt{n} \frac{S_1 - \sigma^2}{\sqrt{2}\sigma^2} \sim N(0, 1)$$

$$\Rightarrow \sqrt{n} (S_1 - \sigma^2) \sim N(0, 2\sigma^2)$$

5) a)  $X \sim \text{Gam}(\alpha, \beta) \quad \alpha, \beta > 0$

$$f(x | \alpha, \beta) = \frac{x^{\alpha-1}}{\Gamma(\alpha) \beta^\alpha} e^{-\frac{x}{\beta}} \cdot \mathbb{1}_{(0, \infty)}(x) = \frac{1}{\Gamma(\alpha) \beta^\alpha} \cdot \mathbb{1}_{(0, \infty)}(x) \cdot \exp\left[(\alpha-1) \ln x - \frac{1}{\beta} x\right]$$

Subst. 2.2.1  
 $\Rightarrow S = s(x_1, \dots, x_n) = \left(\sum_{i=1}^n \ln x_i, \sum_{i=1}^n x_i\right)$  suff. für  $(\alpha, \beta)$ .

b)  $X \sim \text{Be}_2(\alpha, \beta) \quad \alpha, \beta > 0$

$$f(x | \alpha, \beta) = \frac{x^{\alpha-1} (1-x)^{\beta-1}}{B(\alpha, \beta)} \mathbb{1}_{(0,1)}(x) = \frac{1}{B(\alpha, \beta)} \mathbb{1}_{(0,1)}(x) \cdot \exp\left[(\alpha-1) \ln x + (\beta-1) \ln(1-x)\right]$$

$\Rightarrow S = \left(\sum_{i=1}^n \ln x_i, \sum_{i=1}^n \ln(1-x_i)\right)$  suff. für  $(\alpha, \beta)$ .

c)  $X \sim U_{(a,b)} \quad a < b$

$$f(x | a, b) = \frac{1}{b-a} \mathbb{1}_{(a,b)}(x) \exp[0] \quad \Rightarrow \text{keine Exponentialfamilie}$$

$$l(a, b; x_1, \dots, x_n) = \prod_{i=1}^n f(x_i | a, b) = \frac{1}{(b-a)^n} \mathbb{1}_{(a,b)}(x_1) \cdots \mathbb{1}_{(a,b)}(x_n)$$

Sei  $S = s(x_1, \dots, x_n) = (\min x_i, \max x_i)$

$$\Rightarrow l(a, b; x_1, \dots, x_n) = \underbrace{\frac{1}{(b-a)^n} \mathbb{1}_{(a,b)}(\min_{1 \leq i \leq n} x_i) \mathbb{1}_{(a,b)}(\max_{1 \leq i \leq n} x_i)}_{\varphi[S(x_1, \dots, x_n); a, b]} \cdot \underbrace{1}_{\varphi\psi(x_1, \dots, x_n)}$$

6)  $x_1, \dots, x_n$  Stichprobe,  $X \sim N(\mu, \sigma^2)$

ZZ:  $S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x}_n)^2$  best. qu. Verlustfkt.  $L(T; \theta) = (T - \theta)^2$  unzulässig.

d.h.  $\exists$  Schätzfkt.  $T$  mit  $\mathbb{E}(L(T), \sigma^2) \leq \mathbb{E}(L(S_n^2), \sigma^2) \quad \forall \sigma^2 \in (0, \infty)$

o)  $\mathbb{E}(L(S_n^2, \sigma^2)) = \mathbb{E}(S_n^2 - \sigma^2)^2 = \text{Var } S_n^2 = \frac{2\sigma^4}{n-1}$

o) Sei  $T := \frac{n-1}{n} S_n^2$ , d.g.

$$\mathbb{E}(L(\frac{n-1}{n} S_n^2, \sigma^2)) = \mathbb{E}(\frac{n-1}{n} S_n^2 - \sigma^2)^2 = \underbrace{\left(\frac{n-1}{n}\right)^2 \mathbb{E}(S_n^2)^2}_{\left(\frac{2\sigma^4}{n-1} + \sigma^4\right)} - 2 \frac{n-1}{n} \underbrace{\mathbb{E} S_n^2}_{\sigma^2} \sigma^2 + \sigma^4$$

$$= \frac{2\sigma^4}{n} + \frac{\sigma^4}{n^2} < \mathbb{E}(L(S_n^2, \sigma^2)) \quad \forall n \in \mathbb{N}$$

